Xuecan Cui & Jang Schiltz

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Asset Pricing Models with Underlying Time-varying Lévy Processes Bachelier World Congress 2016, New York

> Xuecan CUI Jang SCHILTZ University of Luxembourg

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Time-varying volatility: Empirical studies on the statistical properties of realized and/or implied volatilities have given rise to various stochastic volatility models in the literature, such as the Heston model, CEV models and also stochastic volatility models with jumps etc.

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 Existence of jumps is empirically supported: Carr and Wu (2003), Pan (2002).

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- Time-varying volatility: Empirical studies on the statistical properties of realized and/or implied volatilities have given rise to various stochastic volatility models in the literature, such as the Heston model, CEV models and also stochastic volatility models with jumps etc.
- Existence of jumps is empirically supported: Carr and Wu (2003), Pan (2002).
- Jump intensity is time-varying: Christoffersen et al (2012).

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 Previous studies rely on a specific model structure for volatility and jumps (e.g. Santa-Clara and Yan 2010).

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We introduce a general non-parametric time-varying jump diffusion framework as a natural generalisation of the results from literature (Bollerslev, Todorov and Xu, 2015 JFE).

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- Previous studies rely on a specific model structure for volatility and jumps (e.g. Santa-Clara and Yan 2010).
- We introduce a general non-parametric time-varying jump diffusion framework as a natural generalisation of the results from literature (Bollerslev, Todorov and Xu, 2015 JFE).
- Theoretical part: We assume a time-varying Lévy process, with time-varying drift, volatility and jump intensity parameters, to model the jump diffusion economy. We study an equilibrium asset and option pricing model in this economy.

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- Theoretical part: We assume a time-varying Lévy process, with time-varying drift, volatility and jump intensity parameters, to model the jump diffusion economy. We study an equilibrium asset and option pricing model in this economy.
- Empirical part: Under this general framework, we decompose S&P500 index into time-varying processes of drift, volatility and jump, using the Hodrick-Prescott filter and a particle filter.

Model Build-up

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• An investment of S_t in the *stock market* is governed by:

$$\frac{dS_t}{S_{t^-}} = \mu(t)dt + \sigma(t)dB_t + (e^x - 1)dN_t - \lambda(t)E(e^x - 1)dt, (1)$$

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where S_{t^-} is the value of S_t before a possible jump occurs; $\mu(t)$ and $\sigma(t)$ are the *rate of return* and the *volatility* of the investment.

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where S_{t^-} is the value of S_t before a possible jump occurs; $\mu(t)$ and $\sigma(t)$ are the *rate of return* and the *volatility* of the investment.

The jump part is assumed to be a Poisson process, with jump intensity $\lambda(t)$ and jump size x which follows an arbitrary distribution.

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Money Market Account

We further assume that there is a market for instantaneous borrowing and lending at a *risk-free rate* r(t). The money market account, M_t, follows

$$\frac{dM_t}{M_t} = r(t)dt.$$
 (2)

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The risk-free rate, r(t), will be derived from the general equilibrium later, as a part of the solution.

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Representative Investor

Maximize the expected utility function of life time consumption

$$\max_{c_t} E_t \int_t^T p(t) U(c_t) dt,$$

where c_t is the rate of consumption at time t, U(c) the utility function with U' > 0, U'' < 0, and $p(t) \ge 0$, $0 \le t \le T$ the time preference function.

Assume constant relative risk aversion (CRRA) utility function.

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Assume constant relative risk aversion (CRRA) utility function.

$$U(c) = \left\{egin{array}{cc} rac{c^{1-\gamma}}{1-\gamma} & \gamma > 0, \gamma
eq 1, \ \ln c, & \gamma = 1, \end{array}
ight.$$

where the constant γ is the relative risk aversion coefficient, $\gamma = -c U^{\prime\prime}/U^\prime.$

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Total Wealth

• The total wealth of the representative investor at time *t*:

$$W_t = W_{1t} + W_{2t}$$

where $W_{1t} = \omega W_t$ is invested in the stock market, and $W_{2t} = (1 - \omega)W_t$ is invested in the money market.

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where $W_{1t} = \omega W_t$ is invested in the stock market, and $W_{2t} = (1 - \omega)W_t$ is invested in the money market.

 ω is called the wealth ratio.

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Representative Investor's Optimal Control Problem:

$$\max_{c_t,\omega} E_t \int_t^T p(t) U(c_t) dt, \qquad (3)$$

subject to

$$\begin{aligned} \frac{dW_t}{W_t} &= \omega \frac{dS_t}{S_{t^-}} + (1-\omega) \frac{dM_t}{M_t} - \frac{c_t}{W_t} dt \\ &= [r(t) + \omega \mu(t) - \omega r(t) - \omega \lambda(t) E(e^x - 1) - \frac{c_t}{W_t}] dt + \omega \sigma(t) dB_t \\ &+ \omega (e^x - 1) dN_t, \end{aligned}$$

where $\phi(t) = \mu(t) - r(t)$ is the equity premium.

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where $\phi(t) = \mu(t) - r(t)$ is the equity premium.

Market Clearing: Because there is only one investor in the economy, he has to put all the wealth into the stock market. The general equilibrium occurs at ω = 1, under which the market is cleared.

Equity Premium

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In the production economy with jump diffusion and one representative investor with CRRA utility function, the equilibrium equity premium is given by

 $\phi(t) = \phi_{\sigma}(t) + \phi_{J}(t),$

where $\phi_{\sigma}(t) = \gamma \sigma(t)^2$ -diffusion risk premium

 $\phi_J(t) = \lambda(t) {\sf E}[(1-e^{-\gamma x})(e^x-1)]$ -jump risk premium

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Equity Premium

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 $\phi_J(t) = \lambda(t) E[(1 - e^{-\gamma x})(e^x - 1)]$ -jump risk premium

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• The risk-free rate is a time-varying function:

$$r(t)=\mu(t)-\phi(t)=\mu(t)-\phi_{\sigma}(t)-\phi_{J}(t).$$

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The pricing kernel is given by

$$rac{d\pi_t}{\pi_t} = -r(t)dt - \gamma\sigma(t)dB_t + (e^{y}-1)dN_t - \lambda(t)E(e^{y}-1)dt,$$

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The pricing kernel is given by

 $\frac{d\pi_t}{\pi_t} = -r(t)dt - \gamma\sigma(t)dB_t + (e^y - 1)dN_t - \lambda(t)E(e^y - 1)dt,$

or equivalently, after integration

$$\frac{\pi_T}{\pi_t} = \exp\{-\int_t^T \gamma \sigma(s) dB_s - \int_t^T [r(s) + \frac{1}{2}\gamma^2 \sigma^2(s)] ds - E(e^y - 1) \int_t^T \lambda(s) ds + \sum_{i=1}^{N_t, T} y_i\}.$$

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or equivalently, after integration

$$\frac{\pi_T}{\pi_t} = \exp\{-\int_t^T \gamma \sigma(s) dB_s - \int_t^T [r(s) + \frac{1}{2}\gamma^2 \sigma^2(s)] ds - E(e^y - 1) \int_t^T \lambda(s) ds + \sum_{i=1}^{N_t, T} y_i\}.$$

The random variable y modeling the jump size in the logarithm of the pricing kernel, satisfies E[(e^y − e^{−γx})(e^x − 1)] = 0.

European Call

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The price of a European call, $c(S_t, t)$, in the jump diffusion economy satisfies

$$\frac{\partial c(S_t, t)}{\partial t} + \frac{1}{2}\sigma^2(t)S_t^2\frac{\partial^2 c(S_t, t)}{\partial S^2} + [r(t) - \lambda^Q(t)E^Q(e^x - 1)]S_t\frac{\partial c(S_t, t)}{\partial S} - r(t)c(S_t, t) + \lambda^Q(t)\{E^Q[c(S_te^x, t)] - c(S_t, t)\} = 0,$$

with final condition

$$c(S_T,T) = max(S_T - K,0),$$

where $\lambda^{Q}(t) \equiv \lambda(t)E(e^{y})$: jump intensity in the risk-neutral measure Q, defined by $E^{Q}[f(x)] := \frac{E[e^{y}f(x)]}{E(e^{y})}$, for any function f(x).

European Call

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Pricing formula of a European call option:

$$c(S_t,t) = \sum_{n=0}^{+\infty} e^{-\int_t^T \lambda^Q(s)ds} \frac{(\int_t^T \lambda^Q(s)ds)^n}{n!} E_n^Q [c^{BS}(Se^X e^{-E^Q(e^X-1)\int_t^T \lambda^Q(s)ds},t)]$$

where $c^{BS}(S, t)$ is the Black-Scholes formula price for the European call option and

$$X = \sum_{i=1}^{n} x_i$$

Empirical Part

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Decompose the S&P500 Index into time-varying components, using

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- the Hodrick-Prescott Filter
- a particle filter (Sequential Monte Carlo Method)

Hodrick-Prescott Filter

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Summary

 The Hodrick-Prescott filters was first proposed in Whittaker (1923), then popularized in economics by Hodrick and Prescott (1997).

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¹Ravn and Uhlig (2002)

Hodrick-Prescott Filter

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- The Hodrick-Prescott filters was first proposed in Whittaker (1923), then popularized in economics by Hodrick and Prescott (1997).
- The method serves to decompose the time series
 y_t = ln(S_t) into a trend component τ_t, and a cyclical component c_t:

$$y_t = au_t + c_t$$
, for $t = 1, \dots, T$.

Hodrick-Prescott Filter

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- The method serves to decompose the time series
 y_t = ln(S_t) into a trend component τ_t, and a cyclical component c_t:

$$y_t = au_t + c_t$$
, for $t = 1, \dots, T$.

• Condition: For a given a, τ_t satisfies

$$\min_{\tau} (\sum_{t=1}^{T} (y_t - \tau_t)^2 + a \sum_{t=2}^{T-1} [(\tau_{t+1} - \tau_t) - (\tau_t - \tau_{t-1})]^2),$$

where a = 129600 for monthly data¹.

¹Ravn and Uhlig (2002)

Extract Drift

Data:

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Summary

- S&P500 index, daily, 1985 2014.
- In each month, we use the 5% to 95% quantile of $ln(S_t)$, compute the mean as a monthly data input for HP filter.

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Summary

- S&P500 index, daily, 1985 2014.
- In each month, we use the 5% to 95% quantile of $ln(S_t)$, compute the mean as a monthly data input for HP filter.

As a result, we decompose the stock index into a time-varying trend component T and a component C:

$$ln(S_t) = T + C.$$

 ${\cal T}$ is a monthly drift, ${\cal C}$ the remaining process of volatility plus jumps.

Time-varying Drift



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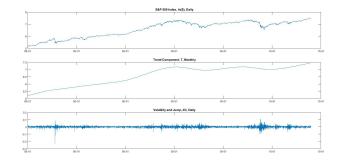


Figure 1.

[mean ($\times 10^{-5}$)	volatility	skewness	kurtosis	
Ì	$\Delta \ln(S)$	31.9	0.0115	-1.3044	31.8	
ĺ	ΔC	1.91	0.0117	-1.2229	30.1	
÷						

Table 1.

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Summary

By taking the difference of the time-varying trend ($\Delta \mathcal{T}),$ we can observe that:

- Regime Switching: Before 2000, stock return was positive. However, after 2000 we can observe that it fluctuates around zero.
- Volatility/Jump Clustering: In negative return periods, there exists jumps and volatility clustering. By contrast, in positive return period, volatility/jump process is much less volatile.

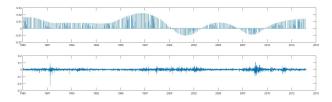


Figure 2.

Filtering Problems

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Summary

- For filtering problem, the data is generated by the state space model, which consists of the observation and state evolution equations,
 - Observation equation: $y_t = f(x_t, \epsilon_t^y)$
 - State evolution: $x_{t+1} = g(x_t, \epsilon_{t+1}^x)$,

where ϵ_{t+1}^{y} is the observation error or "noise", and ϵ_{t+1}^{x} are state shocks.

Filtering Problems

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 - State evolution: $x_{t+1} = g(x_t, \epsilon_{t+1}^x)$,

where ϵ_{t+1}^{y} is the observation error or "noise", and ϵ_{t+1}^{x} are state shocks.

 Particle filters belong to statistical filtering methods, which usually refer to an algorithm for extracting a latent state variable (e.g. volatility) from noisy observations (e.g. stock price/return) using a statistical model.

Particle Filters

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Summary

 Particle filters use a sampling approach with a set of particles to represent the posterior density of a latent state space (Johannes, Polson and Stroud 2009, RFS).

Particle Filters

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Summary

- Particle filters use a sampling approach with a set of particles to represent the posterior density of a latent state space (Johannes, Polson and Stroud 2009, RFS).
- They are simulation-based estimation methods, which include a set of algorithms that estimate the posterior density by directly implementing the Bayesian recursion equations.

Particle Filters

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- Particle filters use a sampling approach with a set of particles to represent the posterior density of a latent state space (Johannes, Polson and Stroud 2009, RFS).
- They are simulation-based estimation methods, which include a set of algorithms that estimate the posterior density by directly implementing the Bayesian recursion equations.
- The state space model used in particle filters can be non-linear and the initial state and noise distributions can take any form required.

SIR Algorithm

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Summary

 The Sampling Importance Resampling (SIR) algorithm is a classical particle filtering algorithm developed by Gordon, Salmond, and Smith (1993).

SIR Algorithm

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Summary

- The Sampling Importance Resampling (SIR) algorithm is a classical particle filtering algorithm developed by Gordon, Salmond, and Smith (1993).
- SIR includes two steps: given samples from $p^N(x_t|y^t)$,
 - S1. Propagation: for i = 1, ..., N, draw $x_{t+1}^{(i)} \sim p(x_{t+1}|x_t^{(i)})$.
 - S2. Resampling: for i = 1, ..., N, draw $z^{(i)} \sim Mult_N(w_{t+1}^{(1)}, ..., w_{t+1}^{(N)})$,

with importance sampling weights $w_{t+1}^{(i)} = \frac{p(y_{t+1}|x_{t+1}^{(i)})}{\sum_{k=0}^{N} p(y_{t+1}|x_{t+1}^{(i)})}$

and set $x_{t+1}^{(i)} = x_{t+1}^{z^{(i)}}$.



State Variable of Volatility

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Summary

As a state space model (for volatility) is necessary to implement particle filters, we assume that following dynamics for the stochastic variance:

$$d
u_t = k(heta -
u_t)dt + \sigma_
u\sqrt{
u_t}dB_t^
u_t$$

where ν_t is a mean-reverting stochastic process. B_t^{ν} is a Brownian motion correlated with B_t , with correlation coefficient ρ .

Filter out Volatility

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 Based on the result from HP filter, we further apply SIR to decompose the time-varying volatility and jump (component C).

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Filter out Volatility

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Summary

- Based on the result from HP filter, we further apply SIR to decompose the time-varying volatility and jump (component C).
- We start with particle filters under stochastic volatility (SV) model without jumps, then we apply particle filters under stochastic volatility and jump (SVJ) model.

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Summary

- Based on the result from HP filter, we further apply SIR to decompose the time-varying volatility and jump (component C).
- We start with particle filters under stochastic volatility (SV) model without jumps, then we apply particle filters under stochastic volatility and jump (SVJ) model.
- The parameters used for the particle filters are taken from Eraker, Johannes and Polson (2003).

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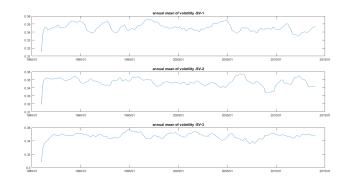
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Summary

- We run particle filters under SV model three times. Estimated volatilities stay around 0.34-0.35. However, the pattern of the volatility processes varies each time.
- Note that the hump shape on the left sides are caused by an adaptation period (around 200 initial data points) needed by the algorithm.



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Summary

Following Eraker, Johannes and Polson (2003), we assume a jump intensity of λ = 0.006, meaning 1 to 2 jumps per year; the jump size follows a normal distribution.

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Summary

- Following Eraker, Johannes and Polson (2003), we assume a jump intensity of λ = 0.006, meaning 1 to 2 jumps per year; the jump size follows a normal distribution.
- With the SVJ model, filtered volatilities decrease to 0.2-0.25, as jumps account for some of the excess variance.

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Summary

- Following Eraker, Johannes and Polson (2003), we assume a jump intensity of λ = 0.006, meaning 1 to 2 jumps per year; the jump size follows a normal distribution.
- With the SVJ model, filtered volatilities decrease to 0.2-0.25, as jumps account for some of the excess variance.
- We detect a high possibility of jumps around 1987-1988, and some other infrequent jumps. Overall jumps are rare in this model. We observe high level of volatilities when the probability of a jump occuring is high.



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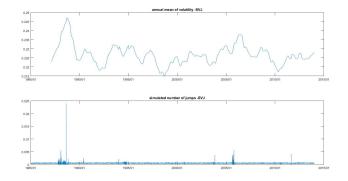


Figure 4. Filtered volatility and jump processes under SVJ model

Future Research

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Summary

The decomposition of the time-varying component of drift, volatility and jumps from S&P500 index using HP filter and particle filter is still a preliminary result.

Future Research

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Summary

- The decomposition of the time-varying component of drift, volatility and jumps from S&P500 index using HP filter and particle filter is still a preliminary result.
- Here we studied the SVJ model only with fixed jump intensity; another possibility is to consider the jump intensity as a time-varying function, for example a function of time-varying drift or volatility, or some other possible exogenous determinant.

Future Research

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Summary

- The decomposition of the time-varying component of drift, volatility and jumps from S&P500 index using HP filter and particle filter is still a preliminary result.
- Here we studied the SVJ model only with fixed jump intensity; another possibility is to consider the jump intensity as a time-varying function, for example a function of time-varying drift or volatility, or some other possible exogenous determinant.
 - It will be interesting to use option data jointly with return data in the filtering methods.

Main References

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Thank you!

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